



Change Points Model

Motivation: Vast amounts of time-ordered, non-stationary data, e.g., stock prices

For $i = 1, 2, \dots, n$,

$$\begin{array}{c} \mathbf{x}_i, \boldsymbol{\beta}^{(l)} \in \mathbb{R}^p \\ y_i = \mathbf{x}_i^\top \boldsymbol{\beta}^{(1)} + \varepsilon_i \\ \downarrow \\ y_i = \mathbf{x}_i^\top \boldsymbol{\beta}^{(2)} + \varepsilon_i \\ \vdots \\ y_i = \mathbf{x}_i^\top \boldsymbol{\beta}^{(L^*)} + \varepsilon_i \end{array} \quad \begin{array}{l} i = \eta_1 \\ \eta := [\eta_1, \dots, \eta_{L^*-1}] \\ L^* \text{ unknown,} \\ \text{given } L^* < L \\ i = \eta_{L^*-1} \end{array}$$

Goal: recover change point locations $\{\eta_l\}_{l=1}^{L^*-1}$ from $\{\mathbf{x}_i, y_i\}_{i=1}^n$, and estimate signals $\{\boldsymbol{\beta}^{(l)}\}_{l=1}^{L^*}$

High dimensional regime: $n, p \rightarrow \infty, n/p \rightarrow \text{fixed constant } \delta$

Existing work

- penalised maximum-likelihood optimisation + partitioning algorithms e.g., [Rinaldo et al. 2021, Xu et al. 2022]
- complementary sketching [Gao and Wang 2022]

Limitations:

- Restricted to certain signal priors (e.g., sparse $\boldsymbol{\beta}^{(l)}$)
- Provide point estimate, without uncertainty quantifications

Our work

Change point inference via Approximate Message Passing (AMP)

Define $\mathbf{B} = [\boldsymbol{\beta}^{(1)} \dots \boldsymbol{\beta}^{(L)}]$, rows of $\mathbf{B} \stackrel{\text{iid}}{\sim} p_{\bar{\mathbf{B}}}$,

$$\boldsymbol{\Theta} := \mathbf{X}\mathbf{B}, \boldsymbol{\Theta} \sim p_{\bar{\boldsymbol{\Theta}}}$$

AMP iteratively produces:

$$\begin{array}{l} \boldsymbol{\Theta}^t = \mathbf{X}\hat{\mathbf{B}}^t - \mathbf{R}^{t-1} (\mathbf{F}^t)^\top \\ \mathbf{R}^t = \mathbf{g}^t(\boldsymbol{\Theta}^t, \mathbf{y}) \end{array}$$

produce residual \mathbf{R}^t
infer change points

$$\begin{array}{l} \mathbf{B}^{t+1} = \mathbf{X}^\top \mathbf{R}^t - \hat{\mathbf{B}}^t (\mathbf{C}^t)^\top \\ \hat{\mathbf{B}}^t = \mathbf{f}^t(\mathbf{B}^t) \end{array}$$

produce signal estimate $\hat{\mathbf{B}}^t$

Theorem (informal):

$$\begin{array}{l} \mathbf{B}^t \stackrel{\mathbb{P}}{\rightarrow} \mathbf{B}\boldsymbol{\nu}_B^t + \mathbf{G}_B^t \quad \text{rows of } \mathbf{G}_B^t \stackrel{\text{iid}}{\sim} N(\mathbf{0}, \boldsymbol{\kappa}_B^t) \\ \boldsymbol{\Theta}^t \stackrel{\mathbb{P}}{\rightarrow} \bar{\boldsymbol{\Theta}}\boldsymbol{\nu}_\Theta^t + \mathbf{G}_\Theta^t \quad \text{rows of } \mathbf{G}_\Theta^t \stackrel{\text{iid}}{\sim} N(\mathbf{0}, \boldsymbol{\kappa}_\Theta^t) \end{array}$$

$\boldsymbol{\nu}_B^t, \boldsymbol{\nu}_\Theta^t, \boldsymbol{\kappa}_B^t, \boldsymbol{\kappa}_\Theta^t$ can be computed deterministically; depend on η

Bayesian approach for choosing denoisers $\mathbf{f}^t, \mathbf{g}^t$

Prior info about change points $\boldsymbol{\eta} \sim p_{\bar{\boldsymbol{\eta}}}$

(Informal): choose \mathbf{f}^t as the Bayes-optimal estimator of signal $\mathbf{B}\boldsymbol{\nu}_B^t$ embedded in Gaussian noise \mathbf{G}_B^t :

$$\mathbf{f}_j^t(\mathbf{B}_j^t) = \mathbb{E} \left[\bar{\mathbf{B}} \mid \bar{\mathbf{B}}\bar{\boldsymbol{\nu}}_B^t + \bar{\mathbf{G}}_{B,j}^t = \mathbf{B}_j^t \right]$$

$$\text{where } \bar{\boldsymbol{\nu}}_B^t := \mathbb{E}_{\bar{\boldsymbol{\eta}}}[\boldsymbol{\nu}_B^t(\bar{\boldsymbol{\eta}})]$$

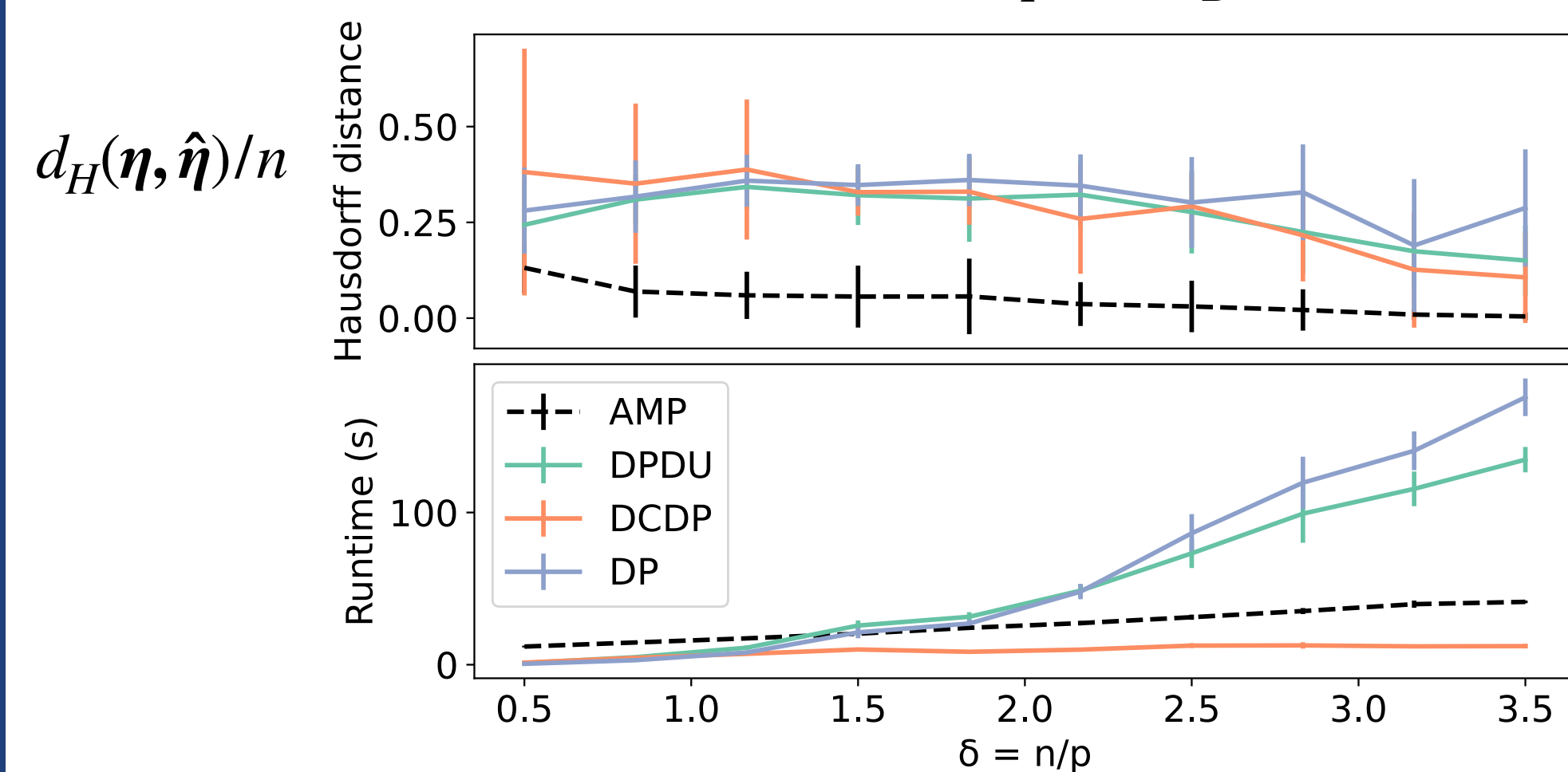
$$\text{rows of } \bar{\mathbf{G}}_B^t \sim N(\mathbf{0}, \bar{\boldsymbol{\kappa}}_B^t), \bar{\boldsymbol{\kappa}}_B^t := \mathbb{E}_{\bar{\boldsymbol{\eta}}}[\boldsymbol{\kappa}_B^t(\bar{\boldsymbol{\eta}})]$$

(Informal): choose \mathbf{g}^t using prior knowledge $\boldsymbol{\eta} \sim p_{\bar{\boldsymbol{\eta}}}$ of change point locations:

$$\mathbf{g}_i^t(\boldsymbol{\Theta}^t, \mathbf{y}) = \mathbb{E} \left[\bar{\boldsymbol{\Theta}}_i \mid \bar{y}_i(\bar{\boldsymbol{\Theta}}_i, \bar{\boldsymbol{\eta}}) = y_i, \bar{\boldsymbol{\Theta}}_i\bar{\boldsymbol{\nu}}_\Theta^t + \bar{\mathbf{G}}_{\Theta,i}^t = \boldsymbol{\Theta}_i^t \right] \cdot \text{other terms}$$

\mathbf{g}_i^t is explicitly dependent on $\bar{\boldsymbol{\eta}}$, row-wise **non-separable** due to sequential structure in $\bar{\boldsymbol{\eta}}$

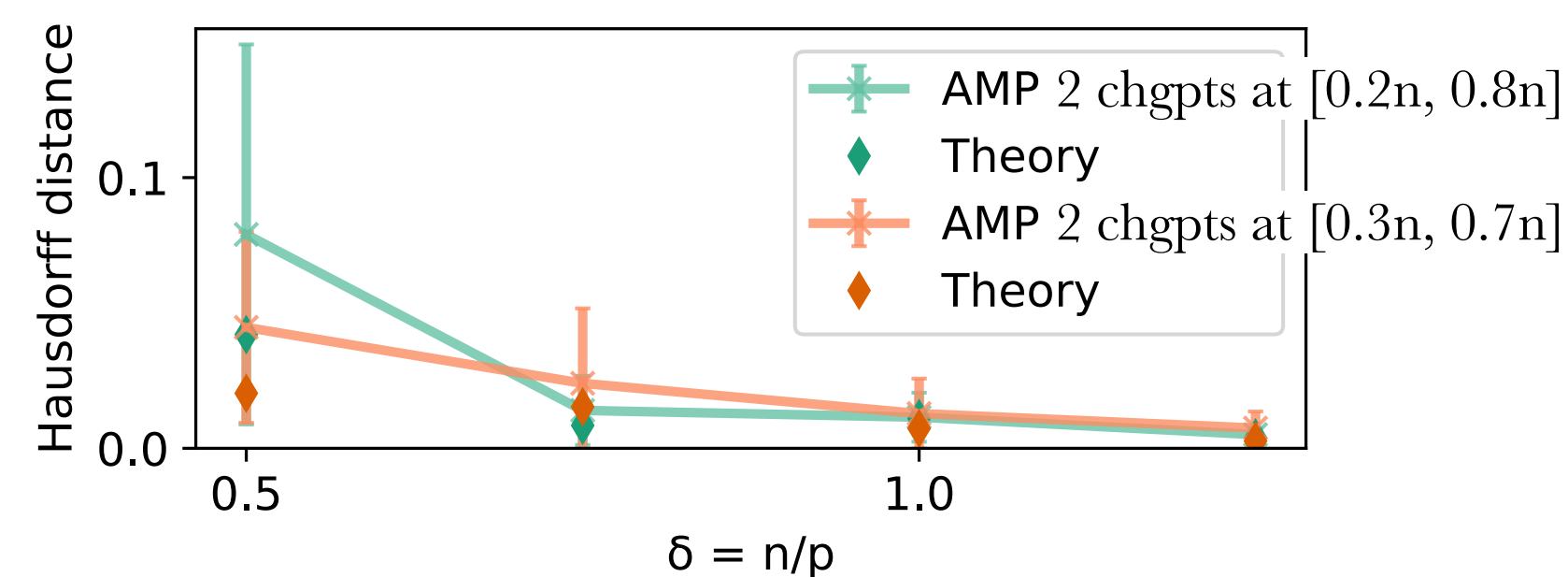
Comparison with existing work sparse $p_{\bar{\mathbf{B}}}, L^* = 3, p = 200$



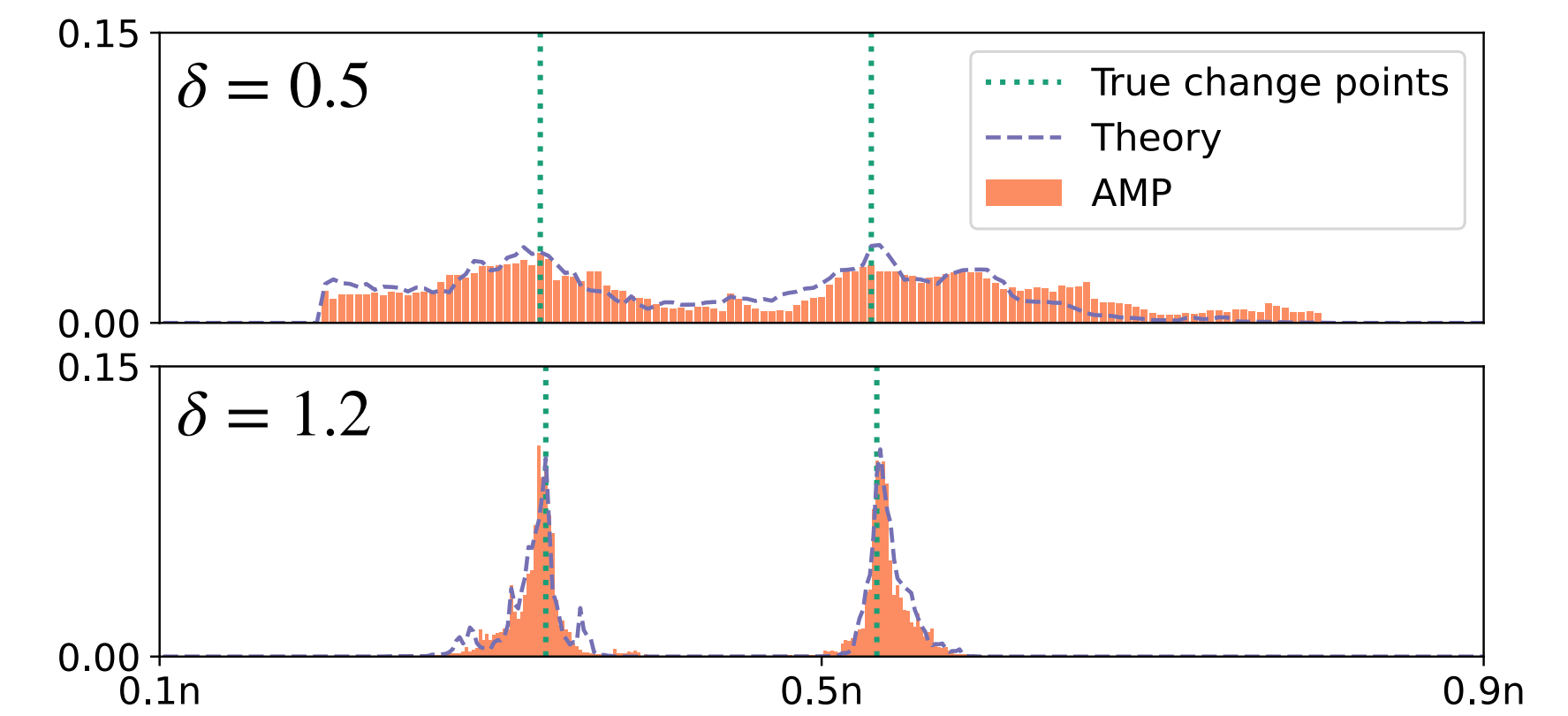
Estimation performance (Hausdorff distance d_H)

Proposition (informal):

$$d_H(\boldsymbol{\eta}, \hat{\boldsymbol{\eta}}(\boldsymbol{\Theta}^t, \mathbf{y})) / n \stackrel{\mathbb{P}}{\rightarrow} \mathbb{E}_{\mathbf{V}_{\bar{\boldsymbol{\Theta}}}, \bar{\boldsymbol{\Theta}}} [d_H(\boldsymbol{\eta}, \hat{\boldsymbol{\eta}}(\bar{\boldsymbol{\Theta}}\boldsymbol{\nu}_\Theta^t + \mathbf{G}_\Theta^t, \bar{y}(\bar{\boldsymbol{\Theta}}, \boldsymbol{\eta})))] / n$$

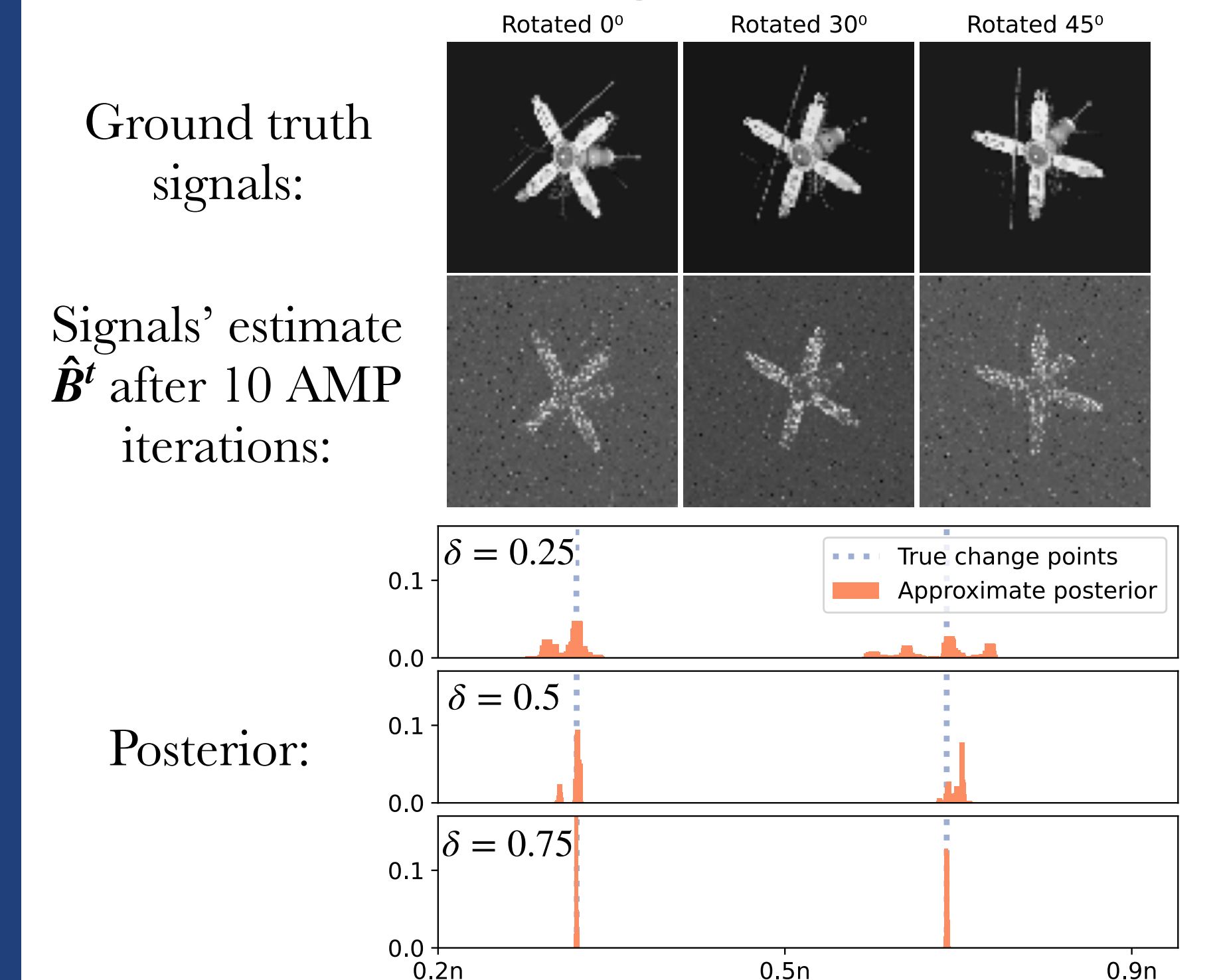


Inference performance (posterior)



$p_{\bar{\mathbf{B}}} = N(\mathbf{0}, \mathbf{I})$, 2 chgpts at $n/3, 8n/15$, $p = 400$
 $L = L^* = 3$, $p_{\bar{\boldsymbol{\eta}}}$ = uniform prior over all configs with chgpts at least $n/5$ apart

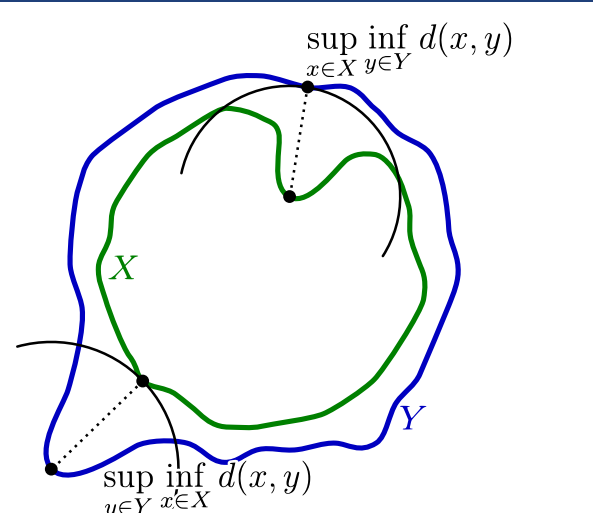
Experiments on real images



Hausdorff distance d_H

Measure of distance between two subsets X, Y of a metric space:

$$d_H(X, Y) := \max \left\{ \sup_{x \in X} d(x, Y), \sup_{y \in Y} d(y, X) \right\}$$



References

- Rinaldo, A., Wang, D., Wen, Q., Willett, R., and Yu, Y. **Localizing Changes in High-Dimensional Regression Models**. In Proceedings of The 24th International Conference on Artificial Intelligence and Statistics, 2021.
- Xu, H., Wang, D., Zhao, Z., and Yu, Y. **Change point inference in high-dimensional regression models under temporal dependence**, 2022. arXiv:2207.12453.
- Gao, F. and Wang, T. **Sparse change detection in high-dimensional linear regression**, 2022. arXiv:2208.06326.